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MANOVA: A METHODOLOGY FOR FORECASTING MANPOWER DEMANDS

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Final Report



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I. Introduction

During an executive session of a Commanders' Conference hosted by the Missile Materiel Readiness Command (MIRCOM) on 13-15 June 1978, the subject of a manpower and work load analysis was discussed. Several commanders suggested to the Chief of Systems Analysis Division, HQ, DARCOM, that an investigative study to examine historical trends in manpower and work load has considerable merit and potential economic benefit. aim at HQ, DAKCOM, is to examine all of their important work generating functions. Most of the functions performed here at TSARCOM have dimensions and are quantifiable. Because these functions are quantifiable, they can be measured, to an acceptable level of confidence, by classical mathematical methodology and analysis. The expectation is amminently optimistic that valid and persuasive relationships can be developed to track trends in work load and associated manpower needs. Where work measurement standards are available, these data will be used in the analysis. In the absence of these kinds of data reports, trend analysis, ratio analysis, linear and multi-variate correlation and regression techniques are among the suggested methodologies to be used. In areas less susceptible to quantitative analysis, professional and mature judgement may be employed through the use of Delphi iterations or other proven schemes.

II. Purpose and Objective

The purpose of MANOVA, a manpower analysis of variance, is to present a comprehensive mathematical methodology designed to facilitate establishment of a statistical baseline for measuring manpower levels associated with work load dependence. Imminently, the objective of this report is preliminary and brief in establishing a functional manpower model. This treatment is subject to expansion provided the need for statistical designs of manpower analysis becomes evolutionary upgraded.

III. Scope and Applicability

The procedures described in this paper are primarily intended to illustrate methods of correlation and regression analysis with special emphasis on the relationship between manpower and work load parameters. The techniques provide useful information for decision making, but the models will not make a decision for the decision maker. In this abbreviated treatment of manpower modeling, several of the statistical parameters which serve to add relevant significance to the various regression equations are included. These statistics include confidence intervals for the regression coefficient or slope of the curve, analysis of variance, and hypothesis testing. This paper will examine the significance of the correlation coefficient and the standard error of the estimate of the independent variable (X, or work load) on the dependent variable (Y, or manpower). The methodology is intended to support and defend the DARCOM Baseline Study for conducting applicable manpower analyses within TSARCOM.

IV. METHODOLOGY

- A. Consideration to the Approach. Operational effectiveness of the US Army Troop Support and Aviation Materiel Readiness Command (TSARCOM) on any functional organization is increased by finding better ways of doing work. The implication in this rather obvious premise is that better ways of doing work are accomplished by managing the use of physical and human resources while minimizing cost or manpower expenditure. Management engineering studies of the type presented herein attempt, through analytical and scientific means, to provide top level management with a sound baseline of manpower vs workload factors whereby quality manpower allocation and economic decisions may be made. During this preliminary phase of the DARCOM Baseline Study, several alternative methods are available to analyze the relationship between manpower expenditure and the quantity of units produced. Some techniques which were suggested by HQ, DARCOM, include:
 - (1) Trend Analysis.
 - (2) Ratio Analysis.
 - (3) Linear Correlation and Regression Analysis.
 - (4) Multi-variate Correlation and Regression Analysis.
 - (5) Other.

Under the category of "Other," the authors have examined the following mathematical and statistical methodologies.

- (1) Exponential Smoothing.
- (2) Bayes' Theorem.
- (3) Bivariate Curve of Best Fit Correlation and Regression, Linear and Curvilinear.

The method proposed by the authors to be superior in measuring relationships between manpower and workload for troop support and aviation system managed items is Simple Regression, Linear and Curvilinear. We support this selection by rationale for eliminating all of the above and setting forth our rationale for selecting Simple Regression, Linear and Curvilinear.

Trend Analysis is essentially a set of observations taken at specified times, usually at equal intervals. The dominate characteristic prevailing in this type of analysis is time dependency. Its strongest attributes are grounded in situations where there exists well defined secular variations, cyclical oscillations, seasonal variations, and random occurances such as acts of nature, strikes, elections, etc. TSARCOM data bases are not compiled in sufficient detail to attempt this kind of manpower analysis. Eliminate this alternative.

Ratio Analysis is very similar in characteristics to
Trend Analysis. In this method, data for each month are
expressed as percentages of monthly trend values. This concept is extremely laborious and inaccurate since too much
of the data are adjusted to approximations, thereby
introducing extraneous errors. Simply stated, if the mean
values of the sample data are not 100%, some arbitrary factor
must be introducted to adjust the data. Eliminate this alternative.

Linear Correlation and Regression Analysis is by far the most superior method of the previous two. However, an assumption must be made in this kind of technique that all data distributions will adjust to a straight line or result in a linear best fit pattern. The assumption associated with this method is too arbitrary. Eliminate this alternative.

Multi-variate Correlation and Regression Analysis is indeed a classical and value-proven statistical scheme. This
concept considers the relationship of three or more independent variables acting on a single dependent variable. For
example, one might attempt to measure the total relationship
of manpower in terms of work units such as the number of
travel vouchers processed, travel orders initiated, card
punch transactions to process travel payment, the number of
completed travel assignments and the number of aborted

pendent though they may seem), the notion that one or more of these "independent" variables is indeed dependent on each the other is very imminent. This concept of interdependency is called autocorrelation. And, while there are techniques to test for and adjust autocorrelation residuals, the technique of multicorrelation and regression is too time consuming and risky for an investigation of magnitude such as the DARCOM Baseline Study. Eliminate this alternative.

Exponential Smoothing is a discrete time series forecasting scheme. In this kind of analysis, the data must
be extremely accurate and some pre-planned data collection
and recording system must be in existence before this procedure
can be accurately manipulated. Also, the derived mathematical equations make use of a "smoothing constant."
This smoothing constant is essentially a variable weighting
factor assigned to selected parameters in the mathematical
equation. Indeed, the smoothing constant is the most
critical variable in the entire scheme. The criteria for
selection of a smoothing constant is fundamentally judgemental
therefore arbitrary. Because the smoothing constant is derived
by heuristic means, the forecast resulting from this methodology

is often viewed with skepticism by high level decision makers.

Mutual fund companies use this methodology to forecast earnings.

Eliminate this alternative.

Bayes Theorem, a probabalistic method, is beyond the scope of consideration although it is perhaps the most precise of all the methods to predict probable occurrences. Its application is highly complex and the TSARCOM data base could not support this kind of classical treatment. It is only mentioned here so that the reader will be aware that the use of the Bayesian property has been considered.

Finally, we reach the method of simple Bivariate Regression, Linear and Curvilinear. This method examines the relationship of one independent variable (work load factors) and one dependent variable (manpower) from historical data, and a curve or model of best fit is generated which in turn describes the data in six equations. For purposes of this paper, we will discuss only four curve fitting equations. They are the:

- (1) Linear or Straight Line Function.
- (2) Exponential Function.
- (3) Parabolic Function.
- (4) Power Function.

For an examination of these four "curve of best fit options" if the data correlates to a prescribed standard (usually a minimum of 70%), we select the equation which best explains the behavior of our data as we have experienced and recorded it. Then, if our history is recorded accurately and stratified in sufficient invervals (say monthly), we arrive at a baseline from which we can extrapolate or project future manpower/workload scenarios. By this method, using selected statistical parameters, we are able to associate levels of confidence in the degree of fit which the mathematical equation describes. While this method is most adaptable to support the TSARCOM data base, there are caveats associated with its use. Some of these caveats are:

- (1) It is a statistical model as opposed to a deterministic model. Therefore, it is not exact and it must be interpreted in the fact.
- (2) Care must be exercised by the user to insure that the input variables do indeed correlate. Quite often, numbers selected at random will show high correlation when in fact they do not have cause-effect relationships. This method is not adequately sensitive to reject this kind of misintended data.

- (3) Interpreting and documenting the results must be clearly annotated.
- (4) There are strict limitations on "how far out" data can accurately be forecast.
- (5) Almost without exception, the object of a statistical study is to furnish a basis for generalization.

 Nevertheless, the strong point is that the proposed method has a high degree of credibility in spite of the caveats of which we are aware. The fact that we are aware of the caveats is the first strong point in our favor. Since 1968, the Air Force has used this method with success in its management engineering procedures. (See AFM 25-5, Management Engineering Procedures, 7 June 1968). US Army Aviation Systems Command, under contractual consulting services with Washington University, St. Louis, MO, used this method with success in a manpower study conducted in 1968 and 1969. The DESCOM Baseline Study of 1978 used the concepts of this methodology. From a practical implementation viewpoint, there are several advantages for TSARCOM in using this methodology. They are:
- (1) Analysts are familiar with the method and have the analytical experience.

- (2) Computer programs are available within the Scientific and Engineering (S&E) computer system (IBM 360/65).
- (3) A computer program is available for partial printout (testing) and graph platting on TSARCOM's minicomputer, the Hewelet Packard 9810.
- (4) The proposed data base is accessible for exercising this model.
- (5) The procedure is easy to explain to high level decision makers.
- (6) There is total in-house capability to support the study logistics.
- (7) The International Mathematical Statistical Library (IMSL), which is available in our Directorate for Management Information Systems, has full documentation and proven test cases for most of the mathematical formulation for this methodology.
- (8) The procedure is upwardly compatible for updating as new data becomes available.
 - (9) Autocorrelation effects are eliminated.
- B. Simple Regression, Linear and Curvilinear. The first step in correlation and regression modeling is logically to select, collect and compile the data. This step results in the construction of a scatter diagram. The scatter diagram will save time by allowing visual elimination of obviously

unacceptable or implausible relationships. The next step is to process the data through a computer program. output will result in a set of parameters by which we can determine which one of four mathematical equations to select for the analysis or to reject all four. This first selection of degree of fit will be based on the standard error of the estimate, the coefficient of correlation and the index of determination. The four mathematical models will be a linear function, an exponential functional, a parabolic function and a power function. General equations describing each of these functions are shown in the Hypothetical Case, paragraph C, below. Once a particular function (equation) is selected as best fit of the data, extrapolating or extending the regression equation is accomplished. There are limitations to which the regression equations are constrained. Mathematical deviations describing each model will not be presented in this paper. Most mathematical statistics textbooks contain the derivation. In general, the best model to select from this methodology is the one with the least standard error of the estimate, the highest index of determination and the highest correlation coefficient. All this being the case, in selecting the best model, management must take into consideration (in addition to the statistical criteria), those judgemental considerations

and real world factors which a mathematical model cannot quantify.

C. The Hypothetical Case. For the purposes of illustration of this methodology, a hypothetical (short) case is presented. We have manpower spaces and workload units produced over a 10 year period. We wish to execute the Simple Regression, Linear and Curvilinear model to determine the best equation for establishing a baseline for projection.

The next pages show the parameters and equations derived from the same data. The computation shows the extrapolation for five projected work load situation years and the number of manpower spaces required to do the forecasted work. Appendices A through F contain details which the more than casual reader might wish to examine.

HYPOTHETICAL DATA BASE

FISCAL YEAR (FY)	MANPOWER (Y)	WORK LOAD (X)	
69	2	17	
70	3	6	
71	4	22	
72	4	37	
73	5	50	
74	6	39	
75	7	31	
76	7	48	
77	7	57	
78	8	40	

Range (X): 6 - 57 Range (Y): 2 - 8

Mean of X = 34.7Mean of Y = 5.3N = 10

Max X = 10 Mzx Y = 60

The four resultant models are:

- (1) Linear, Y = a+bx, Y = 2.1707+0.0902(x) $r^2 = 0.51 r = 0.71 Syx = 1.484 tc = 2.896$
- (2) Exponential, Y = a(exp)(bx), Y = 2.385exp(0.0278x) $r^2 = 0.55 r = 0.74 Syx = 0.318 tc = 3.115$
- (3) Parabolic, $Y = a+bx+cx^2$, $Y = 1.1825+0.1721x-0.0013x^2$ $r^2 = 0.54 \text{ r} = 0.73 \text{ Syx} = 1.546 \text{ tc} = 9.267$
- (4) Power, $Y = ax^b$, $Y = 0.9889x^0.4712$ $r^2 = 0.51 r = 0.71 Syx = 0.332 tc = 2.878$

Graphs of the above curves are shown on pages 17-20.

COMPUTATION SUMMARY

					EXTRAPOLATION			(X)
FUNCTION	r	Syx	tc	<u>60</u>	64	68	72	76
Linear	0.715	1.484	2.896	7	8	8	9	10
Exponential	0.740	0.318	3.115	9	9	10	11	12
Parabolic	0.733	1.546	9.267	7	7	7	7	7
Power	0.713	0.332	2.878	7	8	8	8	8

Select Exponential:

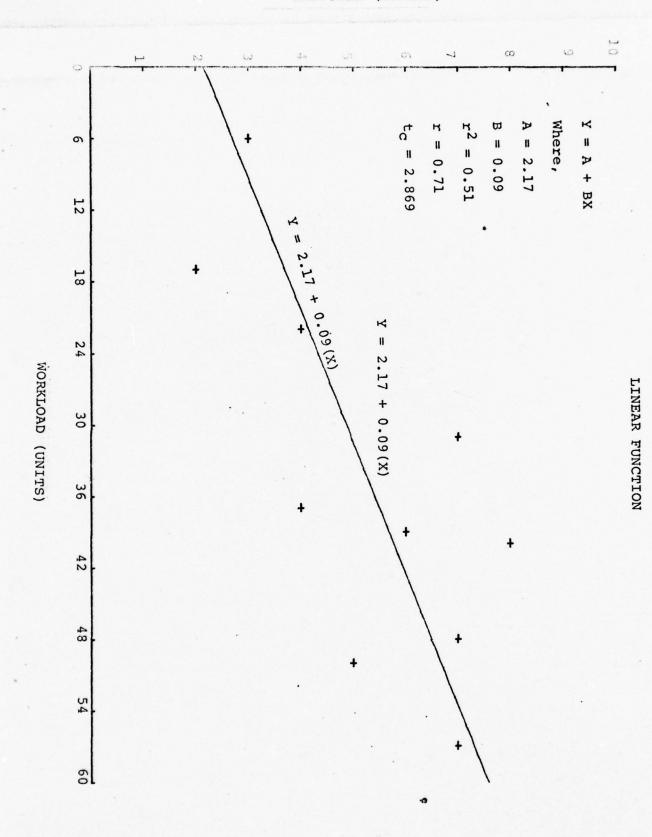
Highest Correlation Coefficient (r) (1)

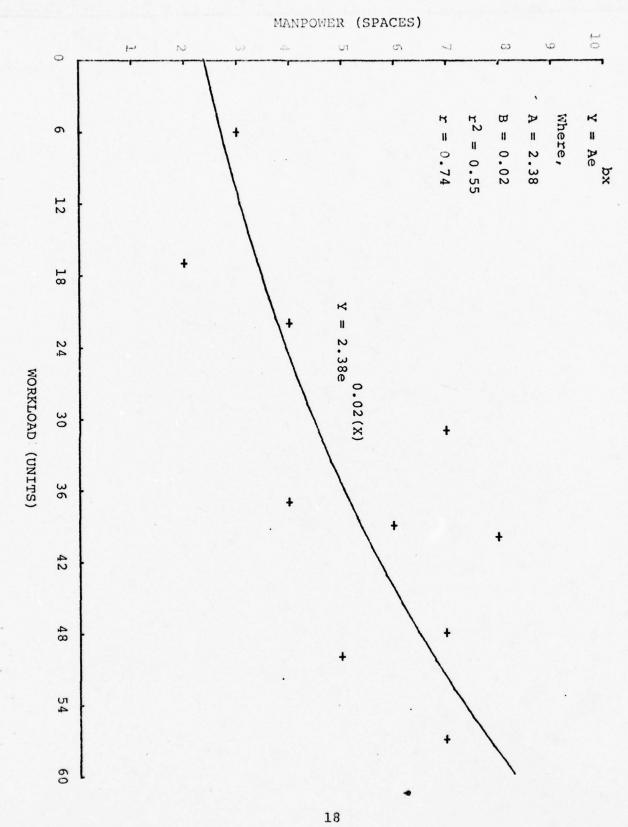
Lowest Standard Error of the Estimate (Syx) (2)

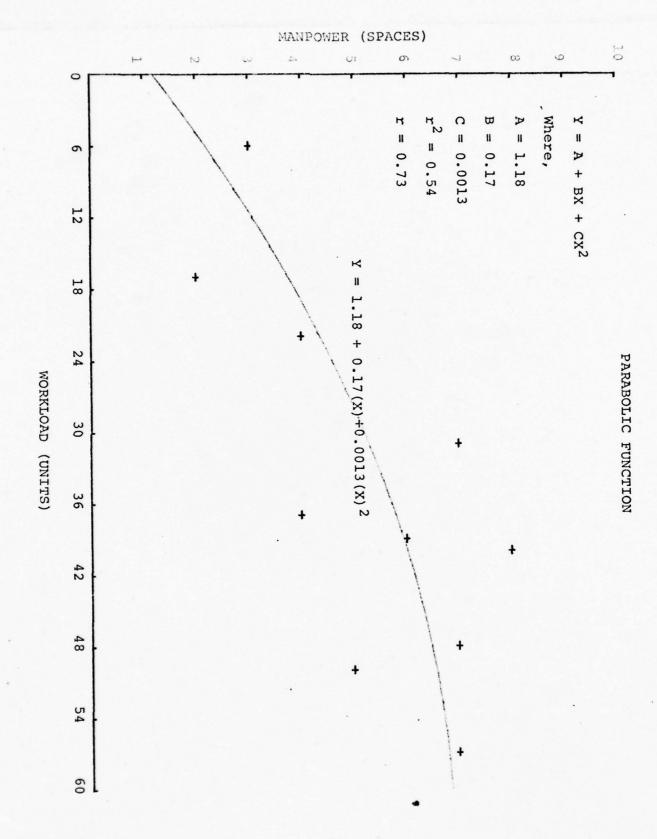
(3)

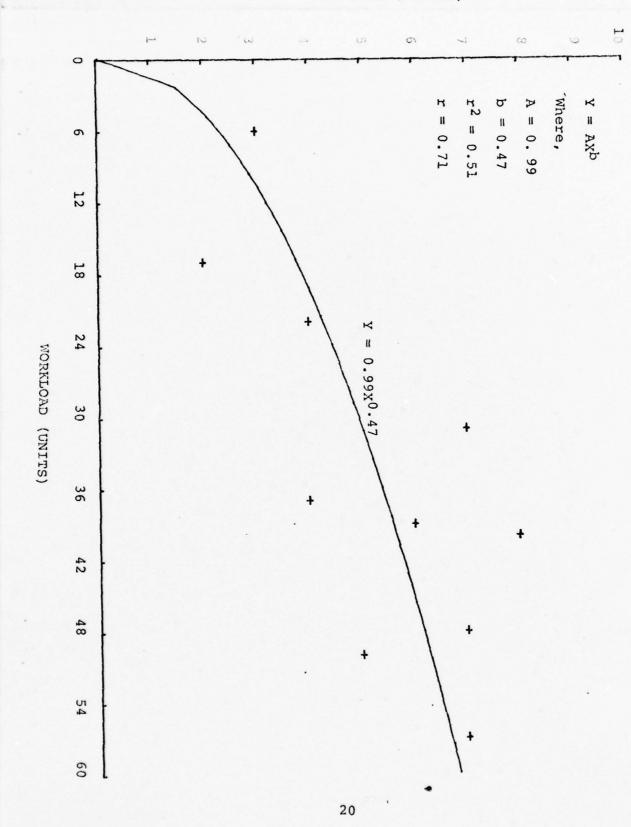
Highest "t" statistic
Interval estimation of slope is smallest (4)(See Appendix D)

The "F" statistic is significant at the 97.5% level (5)









POWER FUNCTION

V. Cost Estimating Considerations. From our hypothetical case, we selected the expoential function to be the curve of best fit for this particular data scenario. Let us now give some consideration to the cost estimates associated with this projected manpower proposition. Our source for making cost projections is DRCCP-ER letter, dated 28 Jul 78, subject: Inflation Guidance (Appendix E).

Let us assume that our starting point is the current fiscal year, FY 78. The largest number of people in our data sample is 8. Let us, for simplicity, assume that these 8 people are GS-11's, step 4. The yearly salary for GS-11/4 is \$20,085. (General Salary Schedule, Oct 77). The total yearly salary for these 8 people is,

8 * \$20,085 = \$160,680

Using OSD/OMB inflation guidance, dated 28 July 1978 (Appendix E), Operations and Maintenance compound indices, our cost projections over the next five fiscal years are shown below:

FY 78 8 * \$20,085 * 1.000 = \$160,680

FY 79 9 * \$20,085 * 1.0630 = \$192,153

FY 80 9 * \$20,085 * 1.1268 = \$203,686

FY 81 10 * \$20,085 * 1.1944 = \$239,895

FY 82 11 * \$20,085 * 1.2613 = \$278,665

FY 83 12 * \$20,085 * 1.3319 = \$321,014

The projected (estimated) salary cost to implement this kind of manpower expansion over the five year period, FY 79 - FY 83 is \$1,235,413.

VI. <u>Summary</u>. We have presented a comprehensive discussion of several methodologies which might apply to the DARCOM Baseline Study with particular emphasis on TSARCOM's capability to exercise the method. We have selected and proposed one method from those alternatives. It is the method of Simple Regression, Linear and Curvilinear. Further, we have illustrated a simple hypothetical case to show, in general, how the method works. Cost considerations have been illustrated using recognized DoD inflation guidance. Through correlation and regression analysis, we believe we have a statistical technique, which, when properly used with accurate data, will provide TSARCOM management an ordered set of baseline alternatives to complement manpower projection decisions.

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APPENDIX A
SUMMATIONS AND GENERAL RESULTS

SUMMATIONS AND GENERAL RESULTS

$$\mathbf{\hat{\Sigma}} X_2 = 347$$

 $\mathbf{\hat{\Sigma}} X^2 = 14,313$
 $\mathbf{\hat{\Sigma}} XY = 2,044$

$$\frac{X}{2}Y = 53$$

 $\frac{X}{2}Y^2 = 317$

$$\frac{N}{X} = 10$$

$$\frac{X}{Y} = 34.7$$

$$= 5.3$$

$$(2x)^2 = 120,409$$

 $(2y)^2 = 2,809$

$$\Sigma (x-\overline{x})^2 = 2272.10$$

$$\mathbf{\xi}(x-\bar{x})^2 = 47.6665$$

$$\mathbf{\Xi}(Y-\overline{Y})^2 = 36.10$$

$$\mathbf{\xi}_{(Y-\overline{Y})^2} = 6.0083$$

APPENDIX B

COMPUTATION OF "t" STATISTIC FOR CORRELATION COEFFICIENTS

COMPUTATION OF "t" STATISTIC FOR CORRELATION COEFFICIENT

I. GENERAL

tc =
$$r - \sqrt{\frac{N-m}{1-r^2}}$$
 $N = 10$
 $m = 2$
 $N-m = 8$

II. PARABOLIC, GENERAL (Compute "F" statistic for correlation coefficient)

$$tc = \frac{r^2}{1-r^2} \cdot \frac{N-m}{m-1}$$

III. LINEAR (SPECIFIC) $r^2 = 0.51185$ r = 0.71544 $1-r^2 = 0.48815$

$$tc = 0.715447 \sqrt{3.0.48815} = 2.896$$

IV. EXPONENTIAL ((SPECIFIC) $r^2 = 0.54814$ r = 0.74036= 0.45186

tc =
$$0.74036 - \sqrt{8 \div 0.45186} = 3.115$$

V. PARABOLIC (SPECIFIC)
$$r^2 = 0.53671$$
 $r = 0.73260$ $1-r^2 = 0.46329$

$$f_{\rm C} = \frac{0.53671}{0.46329} \cdot \frac{8}{1} = \frac{9.267}{1}$$

VI. POWER (SPECIFIC)
$$r^2 = 0.50870$$
 $r = 0.71323$ $1-r^2 = 0.49130$

tc =
$$0.71323 - \sqrt{8 - 0.49130} = 2.878$$

APPENDIX C

COMPUTATION OF "t" STATISTIC FOR REGRESSION COEFFICIENTS

COMPUTATION OF "t" STATISTIC FOR REGRESSION COEFFICIENTS

I. GENERAL

$$t_b = \frac{b}{s_b}$$
 , $s_b = \frac{syx}{\sqrt{2(x-\overline{x})^2}}$

where,

b = Regression Coefficient (Slope)

Syx = Standard Error of the Estimate of X on Y
Sb = Standard Error of the Independent Variable (X)
tb = "t" Distribution Statistic for the
Regression Coefficient

II. LINEAR (SPECIFIC)

$$S_b = 1.48416 - 47.6665 - 0.03114$$

$$t_b = 0.0902 \div 0.03114 = 2.896$$

III. EXPONENTIAL (SPECIFIC)

$$S_b = 0.31792 \div 47.6665 = 0.00667$$

$$t_b = 0.0278 - 0.00667 = 4.168$$

IV PARABOLIC (SPECIFIC)

$$S_b = 1.54571 - 47.6665 = 0.03243$$

$$t_b = 0.1721 - 0.03243 = 5.307$$

$$t_C = -0.0013 \frac{\bullet}{\bullet} 0.03243 = \frac{-0.040}{\bullet}$$

V. POWER (SPECIFIC)

$$S_b = 0.33150 - 47.6665 = 0.00695$$

$$t_b = 0.4712 \div 0.00695 = 67.754$$

APPENDIX D INTERVAL ESTIMATIONS OF REGRESSION EQUATIONS AND HYPOTHESIS TESTING

INTERVAL ESTIMATIONS OF REGRESSION EQUATIONS

AND HYPOTHESIS TESTING

I. GENERAL

A. Hypothesis test for correlation coefficients using the "t" distribution table. ("F" distribution table is used for the Parabolic Function)

 H_O : r = 0 (Null Hypothesis) H_1 : r \neq 0 (Alternative Hypothesis)

B. Hypothesis test for regression coefficients using the "t" distribution table. ("F" distribution table is used for the Parabolic Function)

 H_{0} : b = 0 , H_{0} : c = 0 H_{1} : $b \neq 0$, H_{1} : $c \neq 0$

C. Analysis of Variance (ANOVA) Table for Testing Significance of Regression

Source of Variation	Sum of Squares	Degrees of Freedom	Mean Square	<u>Fo</u>
Regression	ss_R	1 437	MSR	MS _R /MS _t
Error	ss_E	n-2	MSE	
Total	ssy	n-1		

II. LINEAR (SPECIFIC)

A.
$$t_C = 2.896$$
 $t_t = 2.306$ Table Value (8DF, 95% confidence)

 $t_{\rm C}$ > $t_{\rm t}$, Therefore, we reject the null hypothesis, $H_{\rm O}$, and accept the alternative hypothesis with a Type I error, p = 0.05

B. $\approx = 0.05, 1-\%2 = 0.975, t_p (.975,8) = 2.306$

 $t_b = 2.896$

 t_b > t_p , Therefore, we reject the null hypothesis, H_0 , and accept the alternative hypothesis with a Type I error, p = 0.05

Upper "b" limit = 0.162 Lower "b" limit = 0.018 The 95% confidence interval is,

 $0.018 \le b \le 0.162$

In words, the true value of "b" lies in the interval between 0.018 and 0.162, and the statement is made with 95% confidence

C. ANOVA Table (Linear Function)

Regression 14.478 1 18.478 Fo = 8.388 Error 17.622 8 2.203

Total 36.1 9

F(.05,1,8) = 5.32 and we conclude that "b" is not zero at the 95% confidence level

III. EXPONENTIAL (SPECIFIC)

A. $t_C = 3.115$ $t_t = 2.306$ Table Value (8DF, 95% confidence)

 $t_{\rm C}$ > $t_{\rm t}$, Therefore, we reject the null hypothesis, Ho, and accept the alternative hypothesis with a Type I error, p = 0.05

B.
$$\omega \zeta = 0.05$$
, $1-\varepsilon \zeta = 0.975$, t_p (.975,8) = 2.306
 $t_b = 4.168$

Upper "b" limit = 0.0124 Lower "b" limit = 0.0432 The 95% confidence interval is,

 $0.0124 \le b \le .0432$

C. ANOVA Table (Exponential Function)

Regression	0.981	1	0.981	Fo = 9.705
Error	0.808	8	0.101	
Total	1.789	9		

F(.95,1,8) = 5.32 and we conclude that "b" is not zero at the 95% significance level

IV. PARABOLIC (SPECIFIC)

- A. $F_C = 9.267$, $F_t = 4.26$ Table Value (F.95,2,8)
 - F_C > F_t . Therefore, we reject the null hypothesis, H_O , and accept the alternative hypothesis with a Type I error, p = 0.05
- B. eC = 0.05, $F_p(.95,2,8) = 4.26$

 $F_b = 5.307$

 $F_C = ABS(0.040)$

- F_b > F_p , Therefore, we reject the null hypothesis, H_O , and accept the alternative hypothesis with a Type I error, p=0.05
- $F_C \leftarrow F_p$, Therefore, we accept the null hypothesis, H_O , and reject the alternative hypothesis with a Type II error, p = 0.05

Interval estimation is not applicable in this case.

C. ANOVA Table (Parabolic Function)

Regression	19.375	2	9.688	F = 4.055
Error	16.724	7	2.389	
Total	36.099	9		

F(.95,2,8) = 4.26 and we conclude that the regression coefficients "b" and/or "c" could be zero at the 95% significance level

V. POWER (SPECIFIC)

A.
$$t_c = 2.878$$
, $t_t = 2.306$ Table Value (8DF, 95% confidence)

 $t_{\rm C}$ > $t_{\rm t}$, Therefore, we reject the null hypothesis, $H_{\rm O}$, and accept the alternative hypothesis with a Type I error, p = 0.05

B.
$$c = .05$$
, $1 - c = 0.975$ $t_p(.975,8) = 2.306$

 $t_b = 67.754$

t_b > t_p, Therefore, we reject the null hypothesis,
H_o, and accept the alternative hypothesis
with a Type I error, p = 0.05

Upper "b" limit = 0.4872 Lower "b" limit = 0.4552 The 95% confidence interval is,

 $0.4552 \le b \le 0.4872$

C. ANOVA Table (Power Function)

Regression	0.910	1	0.910	F = 8.283
Error	0.879	8	0.109	
Total	1.789	9		

F(.95,1,8) = 5.32 and we conclude that "b" is not zero at the 95% confidence level

APPENDIX E

INFLATION GUIDANCE

DEPARTMENT OF THE ARMY HEADQUARTERS US ARMY MATERIEL DEVELOPMENT AND READINESS COMMAND 5001 EISENHOWER AVE., ALEXANDRIA, VA. 22333

DRCCP-ER

2 8 JUL 1978

SUBJECT: Inflation Guidance

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i. References:

- a. Letter, DRCCP-ER, 28 Dec 77, subject as above.
- b. Letter, DRCCP-ER, 10 Mar 78, subject as above.
- 2. This letter furnishes new inflation indices recently finalized by OSD and DA. These revised indices should be used in the development of the FY 81-85 FOM and the FY 80 Budget. Separate quidance will be issued at a later date regarding use of inflation indices in the preparation of Selected Acquisition Reports (SAR's) and Product Improvement Management Information Reports (PRIMIR's).
- 3. Inclosures 1-8 contain the revised inflation indices. Additional guidance is provided below:
- a. RDTE (Inclosure 1). All major RDTE projects will be inflated. Major RDTE projects are defined as those systems estimated to cost \$75 million or more over the life of the development. Non-major RDTE projects (less than \$75 million) and the in-house salary portion of major RDTE projects will not be inflated for budget estimates. For RDTE in-house salary costs for life cycle cost estimates, the CMA rates will be used. For contract non-major RDTE, the RDTE rates will be used.
- Procurement (Inclosures 2-6). Indices for Aircraft, Mische, Weapons and Tracked Combat Vehicles, Ammunition and Color Procurement Army (OPA) are attached. All Procurement Appropriations will be inflated.

- (1) Procurement of Other than Major Weapons Systems. Indices for the applicable appropriations should be applied to secondary items, modifications and product improvements. Programs not specifically covered by the tables in Inclosures 2-5 should use either the Other Procurement, Army (OPA) indices (Inclosure 6) or the related weapon system indices as appropriate, when specific data is not available. Since items of this type have a short leadtime (typically one year or less), the annual inflation rate rather than the composite rate may be used as appropriate. If the item outlay pattern follows the major item outlay, the composite rate should be used.
- (2) Ammunition (Inclosure 5). Ammunition Production Base Support will be inflated as follows:
- (a) For construction, indices at Inclosure 8 should be used in conjunction with local adjustment factors and guidance contained in AR 415-17, dated 9 Aug 76.
- (b) For equipment procurement, Other Procurement, Army (OPA) indices and outlay rates as contained in Inclosure 6 will be used.
- c. OPA (Inclosure 6). As stated previously in reference la, there is no separate composite index for electronics. All items, including electronic items, in the Other Procurement, Army (OPA) appropriation will use the same index.
- d. OMA (Inclosure 7). As previously stated in reference la, OSD policy in support of Section 806 of the Defense Appropriation Authorization Act of 1977 requires Operations and Maintenance, Army (OMA) to reflect "then year" prices. The indices at Inclosure 7 will be used for that portion of the OMA Appropriation not already covered by special guidance on pay raises, AIF/ASF stabilization increases and POL prices. These indices should also be used for outyear inflation for AIF/ASF type elements of cost/budget estimates not covered by near term special guidance. However, if specific pricing guidance is issued, e.g., revised POL and cost per gallon, costs/estimates must be adjusted accordingly.

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- e. MCA (Inclosure 8). Inclosure 8 contains indices for military construction and new family housing appropriations. These indices should be used in conjuction with local adjustment factors and guidance contained in AR 415-17, dated 9 Aug 76. Additional guidance was contained in reference 1b.
- f. MPA. Since the budgets for the military pay accounts are not inflated, rates for those accounts, which are used only in cost estimating and not program/budget documents, will not be revised at this time. Indices contained at Inclosure 7, reference la, will remain in effect for cost estimating purposes.
- 4. The following additional procedures will be followed in the application of the subject indices:
- a. Cost estimates developed as input to the Army POM and budget must use the attached indices. The only exception will be where an individual program manager has specific contractual arrangements with the prime contractor through contract options or multi-year contracts. The Director of Army Budget must be advised through appropriate inflation guidance channels of all such arrangements before alternate rates are used in POM or budget submissions.
- b. When there is a difference in inflation application between a budget and life cycle cost estimate, the estimate should be submitted both ways to facilitate the cost tracking process.
- c. In developing life cycle cost estimates for Army systems, alternative economic assumptions, e.g., decontrol of gas and oil prices with resultant increase in POL costs to the consumer, may be used so long as the assumptions are clearly defined.
- d. Foreign Military Sales (FMS). Unique rates will not normally be used for FMS. Published rates should be applied to items uniformly whether they are direct Army, FMS, or other customers.

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5. Questions relating to inflation policy should be directed to the HQ DARCOM inflation focal point, Peggy Hombs (DRCCP-ER), AUTOVON 284-9090/9105.

FOR THE COMMANDER:

8 Incl as THOMAS K. HICHTOWER

LTC, GS

Executive Offices

DRCCP-ER SUBJECT: Infla

T: Inflation Guidance

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Harry Diamond Labs, ATTN: Comptroller, Adelphi, MD 20783 US Army Electronics Materiel Readiness Activity, Vint Hill Farms Station, ATTN: DRXEM-CP, Warrenton, VA 22186

US Army Electronics Materiel Readiness Activity, Vint Hill Farms Station, ATTN: DRXEM-NM-M, Warrenton, VA 22186

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US Army Logistics Management Center, ATTN: DRXMC-C-SCAD, Ft Lee, VA 23801

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US Army Management Engineering Training Agency, ATTN:
Comptroller, Rock Island, IL 61201

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DRCCP-F
DRCPP-IF
All Product/Project Managers

OPERATIONS AND MAINTENANCE

						SOME CO	1 1,1014	110,1100 10	DD 0			
1978 1979 1980 AND BEYFIND			BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOUND INDICES COMPOSITE INDICES	ESCALATION RATE		BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979 CO INCIDENT INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE	
. 8300 . 8300	IST YEAR		1.4699	1.5581 1.5759	1.6563 1.6752	1.0560	FY 1987	. 8975 . 8988	.9407	1.0000	1.0710	FY 1978
.1480	2ND YEAR		1.5522 1.5700	1.6454	1.7490 1.7690	1.0560	EX 1088	.9434	1.0000	1.0630	1.0630	FY 1979
			1.6391	1.7375	1.8470 1.8681	1.0560	FY 1989	1.0000	1.0600	1.1268	1.0600	FY 1980
.0120	3RD YEAR	OUTLAY RATES	1.7309 1.7507	1.8348	1.9504	1.0560	FY 1990	1.0600	1.1236	1.1944	1.0600	FY 1981
000 000 000 000	4TH YEAR	ATES	1.8279	1.9375 4.9597	2,0596 2,0832	1.0560	1551 A3	1.1194	1.1865	1.2613	1.0560	FY 1982
0, 99 00 0, 99 00 1, 99 00	STH YEAR		1.9302	2.0460	2.1749 2.1998	1.0560	50 1902	1.1820	1.2530 1.2673	1.3319	1.0560	FY 1983
			2.0383	2.1606	2,2967 2,3230	1.0560	EK : 003	1.2482 1.2625	1.3231	1.4065	1.0560	FY 1984
0.0000	OTH YEAR		2.1525 2.1771	2.2816	2.4254 2.4531	1.0560	76c. As	1.3181	1.3972	1.4853	1.0560	FY 1985
0,0000	7TH YEAR		2.2730 2.2990	2.4094	2.5612	1.0560	265. An	1.3920	1.4755	1.5684	1.0560	FY 1986

MILITARY CONSTRUCTION

SE SE	Se S	7 V	3 1									
OFIGURAS ON OSCI		Tellen,	BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOUND INDICES COMPOSITE INDICES	ESCALATION RATE		BASE YEAR 1980 COMPOUND INDICES COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOUND INDICES COMPOSITE INDICES	ESCALATION RATE	
.0400	IST YEAR		1.5366	1.6441	1.7724	1.0630	Fr 1947	.8670 1.0043	.9276	1.0000	1.0800	FY 1978
. 3600 . 3600	2ND YEAR		1.6334	1.7477	1.8840	1.0630	FY 1988	.9346	1.0000	1.0780 1.2354	1.0780	FY 1979
888	nω		1.7363	1.8578 2.1130	2.0027	1.0630	FY 1989	1.0000	1.0700	1.1535	1.0700	FY 1980
3000	3RD YEAR	OUTLAY RATES	1.8456	: 9748 2.2462	2.1289	1.0630	FY 1990	1.0650	1.1396	1. 2284	1.0650	FY 1981
.1500	ATH YEAR	ATES	1.9619	2.0993	2.2630 2.5739	1.0630	100 m	1.1321	1.2113	1.3058	1.0630	FY 1982
 1000 1000	5TH YEAR		2.0855 2.3721	2. 2315 2. 5381	2.4056 2.7361	1.0630	3001 AE	1.2034	1.2877	1.3881	1.0630	FY 1983
			2.2169	2.3721 2.6980	2.5571	1.0630	ECO 1 A	1.2792	1.3688 1.5568	1.4755	1.0630	FY 1984
25.00 25.00	6TF YEAR		2.3566	2.5215 2.8680	2.7182	1.0630	TY 1004	1.3598	1.4550	1.5685	1.0630	FY 1985
0,000 0,000 0,000	TH YEAR		2.5050 2.8492	2.6804 3.0487	2. 8895 3. 2865	1.0630	FV 1995	1.4455	1.5467	1.6673	1.0630	FY 1986

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1978 1979 1980 AND BEYOND			BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COLUMN INDICES COUPTSITE INDICES	ESCALATION RATE		BASE YEAR 1980 CO (POUND INDICES CO -POSITE INDICES	BASE YEAR 1979 CO APOUND INDICES COMPOSITE INDICES	SASE YEAR 1978 COMPOUND INDICES COMPOSITE INDICES	ESCALATION RATE	
.0700	IST YEAR		1.5900	1.5376	1.8034	1.0540	FY 1987	.8817 .9847	. 9363	1.0000	1.0680	FY 1978
.4300 .4300 .4300	2ND YEAR		1.5260 1.6758	1.6206	1.7308	1.0540	FY 1988	.9416	1.0000	1.0680	1.0680	FY 1979
888			1.6084 3.7663	1.7081 1.8758	1.8243	1.0540	FY 1989	1.0000	1.0620	1.1342	1.0620	FY 1980
.2700 .2700 .2700	3RD YEAR	DUTLAY RATES	1.6952 1.8617	1.8003 1.9771	1.9228	1.0540	FY 1990	1.0560 1.1597	1.1215	1.1977	1.0560	EX 1981
.1500	BYBA HLP	RATES	1.7868 1.9622	1,8976 2.0839	2.0266 2.2256	1.0540	FY 1991	1.1130 1.2223	1.1820	1.2624	1.0540	FY 1982
.0700	5TH YEAR		1. 2833 2.0682	2.0000 2.1961	2.1360 2.3458	1.0540	FY 1992	1.1731	1.2459	1.3306	1.0540	FY 1983
			1.9850 2.1799	2.1080 2.3150	2.2514 2.4725	1.0540	FY 1993	1.2365	1.3131	1.4024	1.0540	FY 1984
200	STH YEAR		2.0921 2.2976	2, 2219 2, 4400	2.3729 2.5060	1.0540	FY 1994	1.3032	1.3840	1.4782	1.0540	FY 1985
0.0000 0.0000	BYEN HILL		2.2051	2.3418 2.5718	2.5011 2.7467	1.0540	FY 1995	1.3736	1.4588	1. 5580 1. 7110	1.0540	FY 1986

1977 1977 1980 AND BEYOND			COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	SASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE		And F YEAR 1980 THE COMPOSITE INDICES COMPOSITE INDICES	8 SE YEAR 1970 COMPOSITE INDICES COMPOSITE INDICES	BACE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	BLAN HULLYNGSS	
.6000	IST YEAR		1.4588 1.4986	1.5507	1.6562	1.0550	[Y 1087	.9101	. 9363	1.0000	:.0700	TY 1975
.3200	2ND YEAR		1.5391	1.6360	1.7473	1.0550	35¢ A3	.9407	1.0000	1.0680	1.0680	FY 1079
			1.6237	1.7260 1.7730	1.8434	1.0550	6361 AS	1.0000	1.0630	1.1353	1.0630	EY 1980
.0700	3RD YEAR	OUTLAY RATES	1.7130	1.8209	1.9447	1.0550	0661 As	1.0580	1.1247	1.2011	1.0580	FY :081
.0100	4TH YEAR	ATES	1.8072	1.9211	2.0517	1.0550	FY 1091	1.1162	1.1865	1.2672	1.0550	FY 1982
0.0000	5TH YEAR		1.90%	2.0267 2.0820	2.1645	1.0550	EV 1002	1.1776	1.2518	1.3369	1.0550	FY 1983
			2.0115	2.1382 2.1965	2.2936	1.0550	FY 1003	1.2423	1.3206	1.4104	1.0550	FY 1984
0.0000 0.0000	6TH YEAR		2. 1221 2. 1799	2.2558 2.3173	2.4092	1.0550	EY 1001	1.3107	1.3932	1.4880	1.0550	FY 1985
0.0000	7TH YEAR		2.2388	2.3799	2.5417	1.0550	. 70	1.3828	1.4699	1.5698	1.0550	FY 1986

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OTHER PROCUREMENT

1973 1979 1980 AND BEYOND			BASE YEAR 1980 COMPOUND INDICES COMPOSITE INDICES	BASE YEAR 1979 COMPOUND INDICES COMPOUND INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE		BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979 CO 42-100 INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE	
.1350 .1350 .1350	IST YEAR		1.4478 1.5841	1.5376 1.6823	1.6421	1.0540	FY 1987	. 2817	.9363	1.0000	1.0680	FY 1978
. 3650 . 3650 . 3650	2ND YEAR		1.5260 1.6696	1.6206 1.7731	1.7308	1.0540	FY 1988	.9416	1.0000	1.0680	1.0680	FY 1979
			1.6084 1.7598	1.7081 1.8689	1.8243	1.0540	FY 1989	1.0000	1.0620	1.1342	1.0620	FY 1980
. 2800 . 2800 . 2800	3RD YEAR	OUTLAY RATES	1.6952 1.8548	1.8003	1.9228	1.0540	FY 1990	1.0560	1.1215 1.2270	1.1977	1.0560	FY 1981
.1500 .1500	4TH YEAR	ATES	1.7868 1.9550	1.8976	2.0266	1.0540	FY 1991	1.1130	1.1820	1.2624	1.0540	FY 1982
.0500	5TH YEAR		1.8833 2.0605	2.0000 2.1883	2. 1360 2. 3371	1.0540	FY 1992	1.1731	1.2459	1.3306 1.4558	1.0540	FY 1983
			1.9850 2.1718	2.1080 2.3064	2.2514	1.0540	FY 1993	1.2365	1.3131	1.4024	1.0540	FY 1984
.0200	6TH YEAR		2.0921	2. 2219 2. 4310	2.3729	1.0540	FY 1994	1.3032	1.3840	1.4782	1.0540	FY 1985
0.0000	7TH YEAR		2.2051	2.3418	2.5011 2.7365	1.0540	FY 1995	1.3736	1.4588	1.5580 1.7046	1.0540	FY 1986

1978 1979 1980 AND BEYOND			SASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1979COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE		BASE YEAR 1980 COMPOSITE INDICES COMPOSITE INDICES	EASE YEAR 1979 COMPOUND INDICES COMPOSITE IMPICES	BASE YEAR 1978 COMPOUNT INDICES COMPOUNT INDICES	ESCALATION RATE	
.1200	IST YEAR		1.4478	1.5376	1.6421	1.0540	FY 1987	.8917	.9363	1.0000	1.0680	FY 1978
.3600 .3600	2ND YEAR		1.5260	1.6206 1.7717	1.7308	1.0540	FY 1988	.9416	1.0000	1.0680	1.0680	FY 1979
			1.6084	1.7081	1.8243	1.0540	FY 1989	1.0000	1.0620	1.1342	1.0620	FY 1980
.3200 .3200 .3200	3RD YEAR	OUTLAY RATES	1.6952 1.8534	1.8003	1.9228	1.0540	FY 1090	1.0560	1. 1215	1.1977	1.0560	FY 1981
.1500	4TH YEAR	ATES	1.7868 1.9534	1.8976	2.02%	1.0540	FY 1991	1.1130 1.2168	1.1820	1.2624	1.0540	FY 1982
.0350 .0350 .0350	5TH YEAR		1.8833	2.0000 2.1866	2.1360 2.3353	1.0540	FY 1992	1.1731	1.2459	1.3306	1.0540	FY 1983
			1.9850 2.1701	2.1080	2.2514	1.0540	FY 1993	1.2365 1.3518	1.3131	1.4024	1.0540	FY 1984
.0150	6TH YEAR		2.0921 2.2873	2. 2219	2.3729	1.0540	FY 1994	1.3032	1.3840	1.4782	1.0540	FY 1985
0.0000	7TH YEAR		2.2051 2.4108	2.3418 2.5603	2.5011 2.7344	1.0540	FY 1095	1.3736	1.4588 1.5949	1.5580 1.7033	1.0540	FY 1986
							F-13					

MEANONS/TRACK COMMAT VEHICLES

1979 1979 1973			BASE YEAR 1980 CO.4POUND INDICES CO.4POSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION RATE		COMPOSITE INDICES COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR ,979 CO "POUND INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE PROJECTS COMPOSITE PROJECTS	ESCALATION BATE	
	ST YEAR		1.4478	1.5376 1.6990	1.6421	1.0540	FY 1987	.8817	. 0363	1.0000	1.0580	87.7.33
.3600 .3600 .3600	2ND YEAR		1.5260 1.6862	1.6206	1.7308	1,0540	FY 1988	.9416	1.0000	1.0680	1.0680	TV 1979
			1.6084 1.7773	1.7081 1.8875	1.8243 2.0158	1.0540	FY 1989	1.0000	1.0620 1.1756	1.1342	1.0620	28 1980
.3400 .3400 .3400	3RD YEAR	TUTLAY RATES	1.6952	1.8003	1.9228	1.3540	FY 1990	1.0560	1.1215	1. 1977 1. 3235	. 0560	EY 1981
1993	4TH YEAR	ATES	1.7868	1.8976	2.0266 2.2394	1.0540	FY 1991	1. 2299	1.1820	1.2624	1.0540	FY 1982
. 0499 0400	5TH YEAR		1.8833 2.0810	2.0000	2.1360 2.3603	1.0540	2661 Ast	i. 1731 i. 2963	1.2459	1. 3306 1. 4703	. 0540	FY 1983
			1.9850 2.1934	2.1080	2.2514	1. 0540	FY 1993	1.2365	1.3131	1.4024	1.0540	FY 1984
.0200	STH YEAR		2.0921	2. 2219	2.3729	1.0540	FY 1994	1.3032	1.3840	1.4782	1.0540	5861 AS
0.0000	TIH YEAR		2.2051 2.4367	2.3418	2.5011 2.7637	:.0540	FY 1995	1.3736	1.4588	1.5580 1.7216	1.0540	FY 1986

1978 1978 1980 AVD 188YOND *			COMPOSITE INDICES	BASE YEAR 1979 COMPOSITE INDICES COMPOSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	ESCALATION PATE		BASE (EAR 1980 COMPOSITE INDICES COMPOSITE INDICES	SUPPLIES ON POSITE INDICES ON POSITE INDICES	BASE YEAR 1978 COMPOSITE INDICES COMPOSITE INDICES	SSCALATION RATE	
888	IST YEAR		1. 4478 1. 5638	1.5376 1.6607	1.5421	1.0540	73 1987	. 8817	.0363	1.0000	1.0680	11974
. 4900 . 4900	2ND YEAR		1.5260 1.6482	1.6206	1.7308	1,0540	FY 1988	.9416	1.0000	1.0680	9.0080	11/10/10
			1.6084 1.7372	1.7081	1.8243	1.0540	FY 1989	1.0000	1.0620	1.1342	1.0620	D801 A
.3000 .3000	3RD YEAR	OUTLAY RATES	1.6952	1.8003	1.9228 2.0758	1.0540	FY 1990	1.0560 1.1406	1.1215	1.1977 1.2937	1.0560	FY 1981
.0500	ATH YEAR	ATES	1.7868	1.8976	2.0256	1.0540	FY 1991	1.1130	1.1820	1.2624	1.0540	FY 1982
.0480 .0430	5TH YEAR		1.8833 2.0341	2. 0000 2. 1603	2.1360 2.3072	1.0540	EX 1093	i.1731 1.2671	1.2459	1. 3306	1.0540	FY 1983
•			1.9850	2.1030 2.2769	2.2514 2.4318	1.0540		1.2365	1.3131	1.4024	1.0540	FY 1984
. 0020 . 0020	STH YEAR		2.0921 2.2598	2.2219	2.3729	1.0540	100 v	1.3032	1.3847	1.4782	1.0540	FY 1985
0.0000	7TH YEAR		2.2051 2.3818	2.3418	2.501; 2.7015	1, 0540	EA 1002	1. 3736	1.45 <i>8</i> 8	1.5580 1.6828	1.0540	FY 1986

APPENDIX F

TABLES

"t" TABLE
"F" TABLE

"	+"	TABLE

			TABLE		
d	t.90	t.95	t.975	٠.99	t.995
1	3.078	6.314	12.706	31.821	63.657
2	1.886	2.920	4.303	6.965	9.925
3	1.638	2.353	3.182	4.541	5.841
4	1.533	2.132	2.776	3.747	4.604
5	1.476	2.615	2.571	3.365	4.032
6 7 8 9	1.449 1.415 1.397 1.333 1.372	1.943 1.895 1.860 1.833 1.813	2.447 2.365 2.306 2.262 2.228	3.143 2.998 2.896 2.821 2.764	3.707 3.499 3.355 3.250 3.169
11	1.363	1.796	2.201	2.718	3.106
12	1.356	1.782	2.179	2.681	3.055
13	1.350	1.771	2.160	2.650	3.012
14	1.365	1.761	2.145	2.624	2.977
15	1.341	1.753	2.131	2.602	2.947
16	1.337	1.746	2.120	2.583	2.921
17	1.333	1.740	2.110	2.567	2.898
18	1.330	1.734	2.101	2.552	2.878
19	1.328	1.729	2.093	2.539	2.861
20	1.325	1.725	2.086	2.528	2.845
21	1.323	1.721	2.080	2.518	2.831
22	1.321	1.717	2.074	2.508	2.819
23	1.319	1.714	2.069	2.500	2.807
24	1.318	1.711	2.064	2.492	2.797
25	1.316	1.703	2.060	2.485	2.787
26	1.315	1.706	2.056	2.479	2.779
27	1.314	1.703	2.052	2.473	2.771
28	1.313	1.701	2.048	2.467	2.763
29	1.311	1.699	2.045	2.462	2.756
30	1.310	1.697	2.045	2.457	2.750
w	1.282	1.645	1.960	2.326	2.576

f(1-ad)

F TABLE

d, = degrees of fivedom for numerator										
1										
12/1	!	2	i .	4	5	0	7	8	9	10
1	161.4	199.5	. 215.7	224.0	230.2	234.0	236.8	238.5	240.5	241.9
· 2	18.51	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.38	19.40
3	10.13	9.55	9.28	9.12	9.01	8.94	8.39	8.85	8.81	8.79
4	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00	5.96
5	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77	4.74
6	5.99	5.14	4.76	4.53	4.39		4.21	4.15	4.10	4.06
7	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68	3.64
8	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39	3.35
10	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18	3.14
10	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02	2.98
11	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90	2.85
12	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80	2.75
13	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71	2.67
14	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65	2.60
15	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59	2.54
16	4.49	3.63	3.24	3.01	2.85	2.74	2.66	2.59	2.54	2.49
17	4.45	3.59	3.20	2.96	2.81	2.70	2.61	2.55	2.49	2.45
18	4.41	3.55	3.16	2.93	2.77	2.66	2.58	2.51	2.46	2.41
19 20	4.38	3.52	3.13	2.90	2.74	2.63	2.54	2.48	2.42	2.38
20	4.35	3.49	3.10	2.87	2.71	2.60	2.51	2.45	2.39	2.35
21	4.32	3.47	3.07	2.84	2.68	2.57			2.37	2.32
22	4.30	3.44	3.05	2.82	2.66	2.55	2.46	2.40	2.34	2.30
23	4.28	3.42	3.03	2.80	2.64	2.53	2.44	2.37	2.32	2.27
24	4.26	3.40	3.01	2.78	2.62	2.51	2.42	2.36	2.30	2.25
25	4.24	3.39	2.99	2.76	2.60	2.49	2.40	2.34	2.28	2.24
26	4.23	3.37	2.98	2.74	2.59	2.47	2.39	2.32	2.27	2.22
21	4.21	3.35	2.96	2.73	2.57	2.46	2.37	2.31	2.25	2.20
28	4.20	3.34	2.95	2.71	2.56	2.45	2.36	2.29	2.24	2.19
29	4.18	3.33	2.93	2.70	2.55	2.43	2.15	2.28	2.22	2.18
30	4.17	3.32	2.92	2.69	2.53	2.42	2.33	2.27	2.21	2.16
40	4.08	3.23	2.84	2.61	2.45	2.34		2.18	2.12	2.08
60	4.00	3.15	2.76	2.53	2.37	2.25	2.17	2.10	2.04	1.99
120	3.92	3.07	2.68	2.45	2.29	2.17	2.09	2.02	1.96	1.91
00	3.84	3.00	2.60	2.37	2.21	2.10	2.01	1.94	1.88	1.83

F(.95.d₁,d₂)